## 1 Lecture: Intro, models

## Repetition of probability notions

## Random variable

- discrete: finite number of realizations (coin, dice, level of service)
- continuous: infinite amount of realizations (car speed, intensity of traffic)

Probability function (discrete variable)

$$
f(x)=P(X=x)
$$

Density function (continuous variable)

$$
P(X \in(a, b))=\int_{a}^{b} f(x) d x
$$

## Examples

- discrete categorical rv

| $f\left(x_{1}, x_{2}\right)$ |  |  |  |
| :---: | :---: | :---: | :---: |
| $x_{1} \backslash x_{2}$ | 1 | 2 | $f\left(x_{2}\right)$ |
| 1 | 0.4 | 0.2 | 0.6 |
| 2 | 0.3 | 0.1 | 0.4 |
| $f\left(x_{1}\right)$ | 0.7 | 0.3 | $=1$ |

- continuous rv

$$
f(x)=a \mathrm{e}^{-a x}, x \geq 0, a>0
$$

$$
P(X>1)=\int_{1}^{\infty} a \mathrm{e}^{-a x} d x=a\left[\frac{-1}{a} \mathrm{e}^{-a x}\right]_{1}^{\infty}=-\left[\mathrm{e}^{-a x}\right]_{1}^{\infty}=\mathrm{e}^{-a}
$$

## Random process

Random variable with time index $x_{t}$. Discrete time $\rightarrow$ random sequence.

$$
x_{1}, x_{2}, x_{3}, \cdots
$$

- discrete random process (with discrete time)

- continuous random process (with continuous time)



## System and its variables



## Bayesian models

Stochastic description of the output $y_{t}$ in dependence on variables in regression vector $\psi_{t}$. The specific relation is given by the parameter $\Theta$. It is in the form of conditional distribution

$$
f\left(y_{t} \mid \psi_{t}, \Theta\right)
$$

## Discrete categorical model

$$
f\left(y_{t} \mid \psi_{t}, \Theta\right)=\Theta_{y_{t} \mid \psi_{t}}
$$

Example for $\psi_{t}=\left[u_{t}, y_{t-1}\right]$, all binary

| $u_{t}$ | 1 | 1 | 2 | 2 |
| :---: | :---: | :---: | :---: | :---: |
| $y_{t-1}$ | 1 | 2 | 1 | 2 |
| $y_{t}=1$ | $\Theta_{1 \mid 11}$ | $\Theta_{1 \mid 12}$ | $\Theta_{1 \mid 21}$ | $\Theta_{1 \mid 22}$ |
| $y_{t}=2$ | $\Theta_{2 \mid 11}$ | $\Theta_{2 \mid 12}$ | $\Theta_{2 \mid 21}$ | $\Theta_{2 \mid 22}$ |$\quad \rightarrow \quad$| $u_{t}$ |
| :---: |
| $y_{t-1}$ | $1_{1}$| 1 | 2 | 1 | 2 |
| :---: | :---: | :---: | :---: | :---: |

For $\left[u_{t}, y_{t-1}\right]=[1,2]$ the prob. of $y_{t}=1$ is 0.9 and $y_{t}=2$ is 0.1 .

Continuous regression model

$$
\begin{gathered}
y_{t}=b_{0} u_{t}+a_{1} y_{t-1}+b_{1} u_{t-1}+\cdots+a_{n} y_{t-n}+b_{n} u_{t-n}+k+e_{t}= \\
=\psi_{t}^{\prime} \theta+e_{t}
\end{gathered}
$$

where noise $e \sim N(0, r), y_{i}$ is output, $u_{t}$ is input and

$$
\begin{gathered}
\psi_{t}=\left[u_{t}, y_{t-1}, u_{t-1}, \cdots, y_{t-n}, u_{t-n}, 1\right]^{\prime} \\
\theta=\left[b_{0}, a_{1}, b_{1}, \cdots a_{n}, b_{n}, k\right]^{\prime} ; \Theta=\{\theta, r\}
\end{gathered}
$$

$n$ is model order.
If $a_{1}=a_{2}=\cdots=a_{n}=0$ the model is static. Otherwise, it is dynamic.
Distribution

$$
f\left(y_{t} \mid \psi_{t}, \Theta\right)=N\left(\psi_{t}^{\prime} \theta, r\right)=\frac{1}{\sqrt{2 \pi r}} \exp \left\{-\frac{1}{2 r}\left(y_{t}-\psi_{t}^{\prime} \theta\right)^{2}\right\}
$$

## State-space model

The state model is

$$
\begin{gathered}
x_{t}=M x_{t-1}+N u_{t}+w_{t} . \\
y_{t}=A x_{t}+B u_{t}+v_{t}
\end{gathered}
$$

Transformation of 2 nd order model regression model to state-space form

$$
y_{t}=b_{0} u_{t}+a_{1} y_{t-1}+b_{1} u_{t-1}+a_{2} y_{t-2}+b_{2} u_{t-2}+k+e_{t}
$$

The state model is

$$
\begin{gathered}
{\left[\begin{array}{c}
y_{t} \\
u_{t} \\
y_{t-1} \\
u_{t-1} \\
1
\end{array}\right]=\left[\begin{array}{ccccc}
a_{1} & b_{1} & a_{2} & b_{2} & k \\
0 & 0 & 0 & 0 & 0 \\
1 & 0 & 0 & 0 & 0 \\
0 & 1 & 0 & 0 & 0 \\
0 & 0 & 0 & 0 & 1
\end{array}\right]\left[\begin{array}{c}
y_{t-1} \\
u_{t-1} \\
y_{t-2} \\
u_{t-2} \\
1
\end{array}\right]+\left[\begin{array}{c}
b_{0} \\
1 \\
0 \\
0 \\
0
\end{array}\right] u_{t}+\left[\begin{array}{c}
e_{t} \\
0 \\
0 \\
0 \\
0
\end{array}\right]} \\
y_{t}=[1,0,0,0,0] x_{t}
\end{gathered}
$$

## Programs

1. T11simCont.sce
simulation of the second order regression model
2. T13simDisc.sce
simulation of discrete model (controlled coin with memory)
$\mathrm{f}(\mathrm{yt}(\mathrm{t}) \mid \mathrm{ut}(\mathrm{t}), \mathrm{yt}(\mathrm{t}-1)), \mathrm{yt}, \mathrm{ut}=1,2$
3. T15simState.sce
simulation with regression model in a state-space form

## 2 Lecture: Estimation

Description of model parameters is given by parameter distribution

$$
f(\Theta \mid d(t))
$$

where $d_{t}=\left\{y_{t}, u_{t}\right\}$ and $d(t)=\left\{d_{0}, d_{1}, d_{2}, \cdots, d_{t}\right\} ; d_{0} \equiv d(0)$ prior data.
Evolution of this distribution is based on the Bayes rule.

## Bayes rule

$$
f(A \mid B, C)=\frac{f(B \mid A, C) f(A \mid C)}{f(B \mid C)} \propto f(B \mid A, C) f(A \mid C)
$$

where
$A$ is what we estimate - parameter $\Theta$
$B$ is what we monitor - new data $y_{t}, u_{t}$
$C$ is old data $d(t-1)$

$$
\underbrace{f(\Theta \mid d(t))}_{\text {posterior pdf }} \propto f\left(y_{t} \mid \psi_{t}, \Theta\right) \underbrace{f(\Theta \mid d(t-1))}_{\text {prior pdf }}
$$

under natural conditions of control

$$
f\left(\Theta \mid u_{t}, d(t-1)\right)=f(\Theta \mid d(t-1))
$$

## Comments

1. Recursion: prior $\rightarrow$ posterior starts with the very prior $f(\Theta \mid d(0))$
2. The computations are recursive - the complexity of parameter distribution must not increase conjugate distribution (Gauss-Wishart, Dirichlet)
3. Recursion on functions - unfeasible. For specific model (categorical, regression) the recursion can be converted to that on statistics, which gives algebraic recursion.
4. Batch estimation (for $t=1,2, \cdots, N$ )

$$
\begin{equation*}
f(\Theta \mid d(N)) \propto \underbrace{f(\Theta \mid d(0))}_{\text {prior pdf }} \underbrace{\prod_{t=1}^{N} f\left(y_{t} \mid \psi_{t}, \Theta\right)}_{\text {likelihood }} \tag{1}
\end{equation*}
$$

5. Results of estimation
(a) posterior distribution $f(\Theta \mid d(t))$
(b) point estimates $\hat{\Theta}_{t}=E[\Theta \mid d(t)]=\int_{\Theta^{*}} \Theta f(\Theta \mid d(t)) d \Theta$

## Estimation of discrete model

Model (for binary $f\left(y_{t} \mid u_{t}, y_{t-1}\right)$ )

| $u_{t}$ | 1 | 1 | 2 | 2 |
| :---: | :---: | :---: | :---: | :---: |
| $y_{t-1}$ | 1 | 2 | 1 | 2 |
| $y_{t}=1$ | 0.2 | 0.9 | 0.5 | 1 |
| $y_{t}=2$ | 0.8 | 0.1 | 0.5 | 0 |

Update - for measured $y_{t}, u_{t}, y_{t-1}$ recompute

$$
S_{y_{t} \mid u_{t}, y_{t-1} ; t}=S_{y_{t} \mid u_{t}, y_{t-1} ; t-1}+1
$$

which means: the combination $\left[y_{t}, u_{t}, y_{t-1}\right]$ has been once more measured.
It is similar to the coin.
Program (est_categ.sce)

## Estimation of regression model

Model

$$
\begin{gathered}
f\left(y_{t} \mid \psi_{t} \Theta\right)=\frac{1}{\sqrt{2 \pi r}} \exp \left\{-\frac{1}{2 r}\left(y_{t}-\psi_{t}^{\prime} \theta\right)^{2}\right\} \propto \\
\propto r^{-0.5} \exp \{-\frac{1}{2}\left[-1, \theta^{\prime}\right] \underbrace{\left[\begin{array}{c}
y_{t} \\
\psi_{t}
\end{array}\right]\left[y_{t}, \psi_{t}\right]}_{D_{t} \text { data matrix }}\left[\begin{array}{c}
-1 \\
\theta
\end{array}\right]\}
\end{gathered}
$$

Statistics

$$
V_{t}, \kappa_{t}
$$

where $V_{t}$ is a square positive definite matrix with the dimension of $D_{t}$, information matrix and $\kappa_{t}$ is a scalar counter of data samples.

Statistics update

$$
\begin{aligned}
V_{t} & =V_{t-1}+D_{t} \\
\kappa_{t} & =\kappa_{t-1}+1
\end{aligned}
$$

Point estimates

$$
\begin{aligned}
& \hat{\theta}=\left(V_{\psi}\right)^{-1} V_{y \psi} \\
& \hat{r}=\frac{V_{y}-\hat{\theta} V_{y \psi}}{\kappa}
\end{aligned}
$$

where $V_{y}=V(1,1), V_{y \psi}=V(2:$ end, 1$), V_{\psi}=V(2:$ end, $2:$ end $)$.
Program (est_regr.sce)

## Batch estimation

According to (1), the estimation can be performed in an off-line mode for the whole measured dataset at once.

Example with

$$
y_{t}=b_{0} u_{t}+a_{1} y_{t-1}+b_{1} u_{t-1}+k+e_{t}
$$

for $t=1,2, \cdots, N$

$$
\begin{gathered}
Y=X \theta+E \\
Y=\left[\begin{array}{c}
y_{1} \\
y_{2} \\
\cdots \\
y_{N}
\end{array}\right], \quad X=\left[\begin{array}{cccc}
u_{1} & y_{0} & u_{0} & 1 \\
u_{2} & y_{1} & u_{1} & 1 \\
& \cdots & & 1 \\
u_{N} & y_{N-1} & u_{N-1} & 1
\end{array}\right] \\
\hat{\theta}=\left(X^{\prime} X\right)^{-1} X^{\prime} Y
\end{gathered}
$$

Program (est_regrBatch.sce)
// Batch estimation of 2 ne order regression model

```
// --------------------------------------------
clc, clear, close, mode(0)
```

```
nd=200; // number of data
```

nd=200; // number of data
r=.1; b0=1, a1=.3, b1=-.6, a2=.3, b2=.1, k=1 // parameters
r=.1; b0=1, a1=.3, b1=-.6, a2=.3, b2=.1, k=1 // parameters
y=zeros(1,nd); // output
y=zeros(1,nd); // output
u=rand(1,nd,'n'); // input
u=rand(1,nd,'n'); // input
// simulation
// simulation
for t=3:nd
for t=3:nd
y(t)=b0*u(t)+a1*y(t-1)+b1*u(t-1)+a2*y(t-2)+b2*u(t-2)+k+sqrt (r)*rand (1, 1,'n');
y(t)=b0*u(t)+a1*y(t-1)+b1*u(t-1)+a2*y(t-2)+b2*u(t-2)+k+sqrt (r)*rand (1, 1,'n');
end
end
// estimation
// estimation
Y=y(3:$)';
X=[u(3:$)' y(2:$-1)' u(2:$-1)' y(1:$-2)' u(1:$-2)' ones(nd-2,1)];
th=inv(X'*X)*X'*Y; // point estimates
b0E=th(1), a1E=th(2), b1E=th(3), a2E=th(4), b2E=th(5), kE=th(6)

```

\section*{Prior information}

Example (coin)
\(x=1,1,2,1,2,2, \cdots\)
1. \(S=[0,0]\)
\begin{tabular}{ccccccc}
\(x\) & 1 & 1 & 2 & 1 & 2 & 2 \\
\(S\) & {\([1,0]\)} & {\([2,0]\)} & {\([2,1]\)} & {\([3,1]\)} & {\([3,2]\)} & {\([3,3]\)} \\
\(\theta\) & {\([1,0]\)} & {\([1,0]\)} & {\(\left[\frac{2}{3}, \frac{1}{3}\right]\)} & {\(\left[\frac{3}{4}, \frac{1}{4}\right]\)} & {\(\left[\frac{3}{5}, \frac{2}{5}\right]\)} & {\(\left[\frac{1}{2}, \frac{1}{2}\right]\)}
\end{tabular}
2. \(S=[10,10]\)
\begin{tabular}{ccccccc}
\(x\) & 1 & 1 & 2 & 1 & 2 & 2 \\
\(S\) & {\([11,10]\)} & {\([12,10]\)} & {\([12,11]\)} & {\([13,11]\)} & {\([13,12]\)} & {\([13,13]\)} \\
\(\theta\) & {\(\left[\frac{11}{21}, \frac{10}{21}\right]\)} & {\(\left[\frac{12}{22}, \frac{10}{22}\right]\)} & {\(\left[\frac{12}{23}, \frac{11}{23}\right]\)} & {\(\left[\frac{13}{24}, \frac{11}{24}\right]\)} & {\(\left[\frac{13}{25}, \frac{12}{25}\right]\)} & {\(\left[\frac{1}{2}, \frac{1}{2}\right]\)}
\end{tabular}

Comparison of estimation without and with prior information



Program (est_init.sce) - try various setting. 0

\section*{Generally to initialization}

Let us have a statistics: \(S_{t}=S_{t-1}+y_{t}\) (sum) and \(\kappa_{t}=\kappa_{t-1}+1\) (count). Let the estimate is \(\hat{\theta}_{t}=S_{t} / \kappa_{t}\). Let our prior knowledge is \(\hat{\theta}_{0}=\theta_{0}\). Then we set:
\[
\kappa_{0}=N, S_{0}=\kappa_{0} \theta_{0}
\]
where \(N\) expresses the strength of the prior information.
Then: \(\hat{\theta}=\left(\kappa_{0} \theta_{0}\right) / \kappa_{0}=\theta_{0}\) and the prior information is obtained as if from \(N\) data records. This is why the several first measured records cannot change it so easy.
and for static regression model

To introduce \(\theta_{0}\), we set
\[
\kappa_{0}=N, V_{0}=\kappa_{0} \begin{array}{|c|ccc|}
\hline 1 & & \theta_{0}^{\prime} & \\
\hline \theta_{0} & 1 & 0 & 0 \\
0 & 1 & 0 \\
& 0 & 0 & 1 \\
\hline
\end{array}
\]

Program (est_init2.sce)

\section*{Programs}

\section*{1. T21estCont_LS.sce}
estimation of 2 nd order regression model
- least squares estimation (off-line)
2. T22estCont_B.sce
estimation of 2 nd order regression model
- Bayesian on-line estimation with statistic update
3. T22estCont_B2.sce
estimation of 2 nd order regression model
- the model for simulation differs from that for estimation
-- Bayesian on-line estimation with statistic update
4. T22estCont_B3.sce
- like the previous one but model order ord can be set
5. T22estCont_B4.sce
estimation of 2 nd order regression model
Estimation with REAL DATA (intensities of traffic in Strahov tunnel)
6. T23estDisc.sce
estimation of discrete model \(\mathrm{f}(\mathrm{y}(\mathrm{t}) \mid \mathrm{u}(\mathrm{t}), \mathrm{y}(\mathrm{t}-1))\) with \(\mathrm{y}, \mathrm{u}\) from \(\{0,1\}\)

\section*{3 Lecture: Prediction}

Estimation of the value of future output.
- predictive pdf
\[
f\left(y_{t+k} \mid y(t-1)\right), k=0,1,2, \cdots
\]
- point prediction
\[
\hat{y}_{t+k}=E\left[y_{t+k} \mid y(t-1)\right]=\int_{y^{*}} y_{t+k} f\left(y_{t+k} \mid y(t-1)\right) d y_{t+k}
\]

Case \(1 k=0\) - output estimation
We are at time \(t, y_{t}\) is not measured, yet and we estimate it on the base of past data.
- model with known parameters
\[
\begin{gathered}
f\left(y_{t} \mid y(t-1)\right)=\text { model } \\
\hat{y}_{t}=\int_{y^{*}} y_{t} f\left(y_{t} \mid y(t-1)\right) d y_{t}
\end{gathered}
\]
- model with unknown parameters
\[
\begin{aligned}
& f\left(y_{t} \mid y(t-1)\right)=\int_{\Theta^{*}} f\left(y_{t}, \Theta \mid y(t-1)\right) d \Theta= \\
= & \int_{\Theta^{*}} \underbrace{f\left(y_{t} \mid y(t-1), \Theta\right)}_{\text {model }} \underbrace{f(\Theta \mid y(t-1))}_{\text {parameter estimate }} d \Theta
\end{aligned}
\]

Case \(2 k>0\) - time prediction (for \(k=1\) )
\[
\begin{gathered}
f\left(y_{t+1} \mid y(t-1)\right)=\int_{y^{*}} \int_{\Theta^{*}} f\left(y_{t+1}, y_{t}, \Theta \mid y(t-1)\right) d \Theta d y_{t}= \\
=\int_{y^{*}} \int_{\Theta^{*}} f\left(y_{t+1} \mid y(t), \Theta\right) f\left(y_{t} \mid y(t-1), \Theta\right) f(\Theta \mid y(t-1)) d \Theta d y_{t}=
\end{gathered}
\]

Point prediction of \(\Theta: \quad f(\Theta \mid y(t-1))=\delta\left(\Theta, \hat{\Theta}_{t-1}\right)\)
\[
=\int_{y^{*}} f\left(y_{t+1} \mid y(t), \hat{\Theta}_{t-1}\right) f\left(y_{t} \mid y(t-1), \hat{\Theta}_{t-1}\right) d y_{t}=
\]
and for \(y_{t}: \quad f\left(y_{t} \mid y(t-1), \hat{\Theta}_{t-1}\right)=\delta\left(y_{t}, \hat{y}_{t}\right)\)
\[
=f\left(y_{t+1} \mid\left[\hat{y}_{t}, y(t-1)\right], \hat{\Theta}_{t-1}\right)
\]

It holds
\[
\int \delta(x, a) f(x) d x=f(a)
\]

\section*{Point prediction with regression model}

The 1st order regression model \(y_{t}=a_{1} y_{t-1}+a_{2} y_{t-2}+b u_{t}+e_{t}\) with known parameters \(a_{1}, a_{2}, b\).
We are at time \(t\) and know all \(u_{t}\), and \(\mathrm{y}(t-1)\).
The prediction is expectation and unknown values are replaced by their predictions (expectations)
\[
\begin{aligned}
y_{t} & =a_{1} y_{t-1}+a_{2} y_{t-2}+b u_{t}+e_{t} \\
\hat{y}_{t} & =a_{1} y_{t-1}+a_{2} y_{t-2}+b u_{t} \\
\hat{y}_{t+1} & =a_{1} \hat{y}_{t}+a_{2} y_{t-1}+b u_{t+1} \\
\hat{y}_{t+2} & =a_{1} \hat{y}_{t+1}+a_{2} \hat{y}_{t}+b u_{t+2}
\end{aligned}
\]

\section*{Full prediction under condition of normality}

Prediction with normal model with known parameters preserves normality. If \(e_{t}\) is normal, all predictions are normal, too.
\[
\begin{aligned}
y_{t} & =a y_{t-1}+b u_{t}+e_{t} \\
y_{t-1} & =a y_{t}+b u_{t+1}+e_{t+1}= \\
& =a\left(a y_{t-1}+b u_{t}+e_{t}\right)+b u_{t+1}+e_{t+1}= \\
& =a^{2} y_{t-1}+a b u_{t}+b u_{t+1}+a e_{t}+e_{t+1} \\
y_{t+2} & =a y_{t+1}+b u_{t+2}+e_{t+2}= \\
& =a^{3} y_{t-1}+a^{2} b u_{t}+a b u_{t+1}+b u_{t+2}+a^{2} e_{t}+a e_{t+1}+e_{t+2}
\end{aligned}
\]
\(\rightarrow\)
\(E\left[y_{t+2} \mid y(t-1)\right]=a^{3} y_{t-1}+a^{2} b u_{t}+a b u_{t+1}+b u_{t+2}\)
\(D\left[y_{t+2} \mid y(t-1)\right]=D\left[a^{2} e_{t}+a e_{t+1}+e_{t+2}\right]=\left(a^{4}+a^{2}+1\right) r\)
Predictive pdf
\[
f\left(y_{t+2} \mid y(t-1)\right)=N_{y_{t+2}}\left(E\left[y_{t+2} \mid y(t-1)\right], D\left[y_{t+2} \mid y(t-1)\right]\right)
\]

\section*{Programs}

\section*{1. T31preCont.sce}
np-step prediction with continuous model (known parameters)
2. T32preCont_Adapt.sce
n-step prediction with continuous model (with estimation)
3. T32preCont_Adapt2.sce
n-step prediction with continuous model (with estimation)
- the model for simulation differs from that for estimation
4. T32preCont_Adapt3.sce
np-step prediction with continuous model (with estimation)
- real data (intensity) from Strahov tunnel are used
5. T33preCat_Off.sce
prediction with discrete model (off-line), known parameters
6. T34preCat_OffEst.sce
prediction with discrete model (off-line), unknown parameters
7. T35preCat_OnEst.sce
prediction with discrete model (on-line)

\section*{4 Lecture: State-space model}

Model
\[
\begin{array}{ll}
f\left(x_{t} \mid x_{t-1}, u_{t}\right) & \text { model of the state } \\
f\left(y_{t} \mid x_{t}, u_{t}\right) & \text { model of the output }
\end{array}
\]
is generated by the equations
\[
\begin{aligned}
x_{t} & =M x_{t-1}+N u_{t}+w_{t} \\
y_{t} & =A x_{t}+B u_{t}+v_{t}
\end{aligned}
\]
where \(M, N, A, B\) are matrices, \(w_{t}\) and \(v_{t}\) white noises with covariance matrices \(r_{w}\) and \(r_{v}\).

\section*{Estimation}

\section*{State description}
\[
f\left(x_{t-1} \mid d(t-1)\right) \underbrace{\rightarrow}_{\text {prediction }} f\left(x_{t} \mid d(t-1)\right) \underbrace{\rightarrow}_{\text {filtration }} f\left(x_{t} \mid d(t)\right)
\]

\section*{Evolution}
\[
\begin{aligned}
f\left(x_{t} \mid d(t-1)\right) & =\int_{x_{t-1}^{*}} f\left(x_{t} \mid x_{t-1}, u_{t}\right) f\left(x_{t-1} \mid d(t-1)\right) \text { prediction } \\
f(\underbrace{x_{t}}_{\Theta} \mid d(t)) & \propto \underbrace{f\left(y_{t} \mid x_{t}, u_{t}\right)}_{\text {model }} f(\underbrace{x_{t}}_{\Theta} \mid d(t-1)) \text { Bayes }
\end{aligned}
\]
! In the above derivation Natural Conditions of Control are used !

\section*{Kalman filter}

For normal model and normal prior
Notation
\[
\begin{aligned}
f\left(x_{t} \mid x_{t-1}, u_{t}\right) & =N_{x_{t}}\left(M x_{t-1}+N u_{t}, r_{w}\right) \\
f\left(y_{t} \mid x_{t}, u_{t}\right) & =N_{y_{t}}\left(A x_{t}+B u_{t}, r_{v}\right)
\end{aligned}
\]
and
\[
\begin{aligned}
f\left(x_{t-1} \mid d(t-1)\right) & =N_{x_{t-1}}\left(x_{t-1 \mid t-1}, R_{t-1 \mid t-1}\right) \\
f\left(x_{t} \mid d(t-1)\right) & =N_{x_{t}}\left(x_{t \mid t}, R_{t \mid t}\right) \\
f\left(x_{t} \mid d(t)\right) & =N_{x_{t}}\left(x_{t \mid t}, R_{t \mid t}\right)
\end{aligned}
\]

Substitution into the evolution equations gives Kalman filter (KF)
\[
\begin{aligned}
& \text { Kalman filter } \\
& x_{t \mid t-1}=M x_{t-1 \mid t-1}+N u_{t} \\
& R_{t \mid t-1}=r_{x}+M R_{t-1 \mid t-1} M^{\prime} \\
& y_{p}=A x_{t \mid t-1}+B u_{t} \\
& R_{p}=r_{y}+A R_{t \mid t-1} A^{\prime} \\
& R_{t \mid t}=R_{t \mid t-1}-R_{t \mid t-1} A^{\prime} R_{p}^{-1} A R_{t \mid t-1} \\
& K=R_{t \mid t} A^{\prime} r_{y}^{-1} \\
& x_{t \mid t}=x_{t \mid t-1}+K\left(y_{t}-y_{p}\right)
\end{aligned}
\]
\[
x_{t \mid t-1}=M x_{t-1 \mid t-1}+N u_{t} \quad \text { state prediction }
\]
\[
y_{p}=A x_{t \mid t-1}+B u_{t} \quad \text { output prediction }
\]

Kalman gain
state correction

\section*{Nonlinear model}
\[
\begin{aligned}
x_{t} & =g\left(x_{t-1}, u_{t}\right)+w_{t} \\
y_{t} & =h\left(x_{t}, u_{t}\right)+v_{t}
\end{aligned}
\]

\section*{Example}

For
\[
x_{t}=\left[\begin{array}{l}
x_{1} \\
x_{2}
\end{array}\right]_{t}, \quad u_{t}, y_{t}
\]
the model is
\[
\begin{aligned}
x_{1 ; t} & =\exp \left\{-x_{1 ; t-1}-x_{2 ; t-1}\right\}+u_{t}+w_{t} \\
x_{2 ; t} & =x_{1 ; t-1}-0.3 u_{t}+w_{2 ; t} \\
y_{t} & =x_{2 ; t}+v_{t}
\end{aligned}
\]

\section*{Linearization}

Is done using first two terms of Taylor expansion of nonlinear functions at the point of last point estimate. For the state equation it is \(\hat{x}_{t-1}\) and for the output equation it is \(\hat{x}_{t}\).

Generally, i.e. for a general value \(x\) the expansion reads
\[
\begin{gathered}
g\left(x, u_{t}\right) \doteq g\left(\hat{x}_{t-1}, u_{t}\right)+g^{\prime}\left(\hat{x}_{t-1}, u_{t}\right)\left(x-\hat{x}_{t-1}\right) \\
\quad h\left(x, u_{t}\right) \doteq h\left(\hat{x}_{t}, u_{t}\right)+h^{\prime}\left(\hat{x}_{t}, u_{t}\right)\left(x-\hat{x}_{t}\right)
\end{gathered}
\]

\section*{Remarks}
1. \(x_{t}\) and \(x_{t-1}\) are random variables. \(x\) is their general value, \(\hat{x}_{t}\) and \(\hat{x}_{t-1}\) are special values: \(\hat{x}_{t}\) is the point estimate of \(x_{t}\) and \(\hat{x}_{t-1}\) is point estimate of \(x_{t-1}\).
2. Linearization can be applied only to nonlinear parts of the model. The linear parts can stay as they are.

The derivatives \(g^{\prime}\) and \(h^{\prime}\) are
\[
g^{\prime}\left(\hat{x}_{t-1}, u_{t}\right)=\left[\begin{array}{cccc}
\frac{\partial g_{1}}{\partial x_{1}} & \frac{\partial g_{1}}{\partial x_{2}} & \cdots & \frac{\partial g_{1}}{\partial x_{n}} \\
\cdots & \cdots & \cdots & \cdots \\
& & \cdots & \\
\frac{\partial g_{n}}{\partial x_{1}} & & \cdots & \frac{\partial g_{n}}{\partial x_{n}}
\end{array}\right]_{\mid x=\hat{x}_{t-1}} \quad, \quad h^{\prime}\left(\hat{x}_{t}, u_{t}\right)=\left[\begin{array}{cccc}
\frac{\partial h_{1}}{\partial x_{1}} & \frac{\partial h_{1}}{\partial x_{2}} & \cdots & \frac{\partial h_{1}}{\partial x_{n}} \\
\cdots & \cdots & \cdots & \cdots \\
& & \cdots & \\
\frac{\partial h_{m}}{\partial x_{1}} & & \cdots & \frac{\partial h_{m}}{\partial x_{n}}
\end{array}\right]_{\mid x=\hat{x}_{t}}
\]

After substitution the linearization into the model, we have (for \(x=x_{t-1}\) in the case of the state
equation and \(x=x_{t}\) for output equation) we obtain the linearized model
\[
\begin{aligned}
x_{t} & =\bar{M} x_{t-1}+F+w_{t} \\
y_{t} & =\bar{A} x_{t}+G+v_{t}
\end{aligned}
\]
where
\[
\begin{gathered}
\bar{M}=g^{\prime}\left(\hat{x}_{t-1}, u_{t}\right), \quad F=g\left(\hat{x}_{t-1}, u_{t}\right)-\underbrace{g^{\prime}\left(\hat{x}_{t-1}, u_{t}\right)}_{\bar{M}} \hat{x}_{t-1}, \\
\bar{A}=h^{\prime}\left(\hat{x}_{t}, u_{t}\right), \quad G=h\left(\hat{x}_{t}, u_{t}\right)-\underbrace{h^{\prime}\left(\hat{x}_{t}, u_{t}\right)}_{\bar{A}} \hat{x}_{t} .
\end{gathered}
\]

Example (continuation) - \(\cdots\) only first equation is nonlinear
\[
\begin{gathered}
g_{1}\left(x, u_{t}\right)=\exp \left\{-x_{1}-x_{2}\right\}+u_{t} \\
g_{2}\left(x, u_{t}\right)=x_{1}-0.3 u_{t} \\
g_{1}^{\prime}\left(x, u_{t}\right)=\left[\frac{\partial g_{1}}{\partial x_{1}}, \frac{\partial g_{1}}{\partial x_{2}}\right]=\left[-\exp \left\{-x_{1}-x_{2}\right\},-\exp \left\{-x_{1}-x_{2}\right\}\right]
\end{gathered}
\]
\[
\begin{gathered}
g_{2}^{\prime}\left(x, u_{t}\right)=\left[\frac{\partial g_{2}}{\partial x_{1}}, \frac{\partial g_{2}}{\partial x_{2}}\right]=[1,0] \\
\bar{M}=\left[\begin{array}{cc}
-\exp \left\{-x_{1}-x_{2}\right\}, & -\exp \left\{-x_{1}-x_{2}\right\} \\
1
\end{array}\right] \\
F=\left[\begin{array}{c}
\exp \left\{-x_{1}-x_{2}\right\}+u_{t} \\
x_{1}-0.3 u_{t}
\end{array}\right]-\bar{M} x_{t-1}
\end{gathered}
\]

The output equation is linear with \(\bar{A}=[0,1]\)
Fully linearized model is
\[
\begin{aligned}
x_{t} & =\bar{M} x_{t-1}+F+w_{t} \\
y_{t} & =\bar{A} x_{t}+v_{t}
\end{aligned}
\]

With
\[
N=\left[\begin{array}{l}
0 \\
0
\end{array}\right], \quad G=0, \quad B=0
\]
we can use subroutine Kalman
\[
[\mathrm{xt}, \mathrm{Rx}, \mathrm{yp}]=\mathrm{Kalman}(\mathrm{xt}, \mathrm{yt}, \mathrm{ut}, \bar{M}, \mathrm{~N}, \mathrm{~F}, \bar{A}, \mathrm{~B}, \mathrm{G}, \mathrm{Rw}, \mathrm{Rv}, \mathrm{Rx})
\]

\section*{Programs}

\section*{1. T46statEst_KF.sce}
state estimation (Kalman filter)
2. T47statEst_Noise.sce

Kalman as a noise filter
3. T48statEst_NL.sce
nonlinear model estimation (T48statEst_L.sce - linear version)
4. T48statEst_Par.sce
unknown parameters

\section*{5 Lecture: Control}

Minimum variance control - in each step \(t\) minimizes \(E\left[y^{2}\right]\).
Model (e.g. first order)
\[
y_{t}=b_{0} u_{t}+a_{1} y_{t-1}+b_{1} u_{t-1}+k+e_{t}
\]
\(E\left[y_{t}^{2}\right]=\left(b_{0} u_{t}+a_{1} y_{t-1}+b_{1} u_{t-1}+k\right)^{2}+r \rightarrow \min\)
\(\rightarrow b_{0} u_{t}+a_{1} y_{t-1}+b_{1} u_{t-1}+k=0\)
\[
u_{t}=-\frac{1}{b_{0}}\left(a_{1} y_{t-1}+b_{1} u_{t-1}+k\right)
\]

Often unstable !!!

\section*{Derivation of optimal control}

Model
\[
y_{t}=\psi_{t}^{\prime} \theta+e_{t}
\]

Criterion
\[
J=E\left[\sum_{t=1}^{N} J_{t} \mid d(0)\right]
\]
where \(J_{t}=y_{t}^{2}+\omega u_{t}^{2}\).
Bellman equations
\[
\begin{gathered}
\varphi_{t}=E\left[\varphi_{t+1}^{*}+J_{t} \mid u_{t}, d(t-1)\right] \quad \text { expectation } \\
\varphi_{t}^{*}=\min _{u_{t}} \varphi_{t} \quad \text { minimization }
\end{gathered}
\]
for \(t=N, N-1, N-2, \cdots, 1\).

\section*{Control with regression model}

Regression model in state-space form (2nd order)
\[
x_{t}=M x_{t-1}+N u_{t}+w_{t}
\]
where \(x_{t}=\left[y_{t}, u_{t}, y_{t-1}, u_{t-1}, \cdots y_{t-n+1}, u_{t-n+1}\right]^{\prime}\).
The penalty can be written as
\[
\begin{equation*}
y_{t}^{2}+\omega u_{t}^{2}=x_{t}^{\prime} \Omega x_{t} \tag{2}
\end{equation*}
\]
where \(\Omega\) is a diagonal matrix
\[
\Omega=\left[\begin{array}{lllll}
1 & & & & \\
& \omega & & & \\
& & 0 & & \\
& & & \cdots & \\
& & & & 0
\end{array}\right]
\]

Bellman equations, where we guess the form of \(\varphi_{t+1}^{*}=x_{t}^{\prime} R_{t+1} x_{t}\)
\[
\begin{gathered}
E\left[x_{t}^{\prime} R_{t+1} x_{t}+x_{t}^{\prime} \Omega x_{t} \mid u_{t}, d(t-1)\right]=E\left[x_{t}^{\prime} U x_{t} \mid u_{t}, d(t-1)\right]= \\
=\left(M x_{t-1}+N u_{t}\right)^{\prime} U\left(M x_{t-1}+N u_{t}\right)+\rho= \\
=x_{t-1}^{\prime} \underbrace{M^{\prime} U M}_{C} x_{t-1}+2 u_{t}^{\prime} \underbrace{N^{\prime} U M}_{B} x_{t-1}+u_{t}^{\prime} \underbrace{N^{\prime} U N}_{A} u_{t}+\rho= \\
=u_{t}^{\prime} A u_{t}+2 u_{t}^{\prime} A \underbrace{A^{-1} B}_{S_{t}} x_{t-1}+x_{t-1}^{\prime} S_{t}^{\prime} A S_{t} x_{t-1}+ \\
\quad+\underbrace{x_{t-1}^{\prime} C x_{t-1}-x_{t-1}^{\prime} S_{t}^{\prime} A S_{t} x_{t-1}}_{x_{t-1} R_{t} x_{t-1}}+\rho= \\
=\left(u_{t}+S_{t} x_{t-1}\right)^{\prime} A\left(u_{t}+S_{t} x_{t-1}\right)+x_{t-1}^{\prime} R_{t} x_{t-1}+\rho
\end{gathered}
\]

Optimal \(u_{t}=S_{t} x_{t-1}\).

\section*{Recursion}

\section*{Optimization}
\[
\begin{aligned}
R_{N+1} & =0 \\
\text { for } t & =N, N-1, \cdots, 1 \\
U & =R_{t+1}+\Omega \\
A & =N^{\prime} U N \\
B & =N^{\prime} U M \\
C & =M^{\prime} U M \\
S_{t} & =A^{-1} B \\
R_{t} & =C-S_{t}^{\prime} Q S_{t}
\end{aligned}
\]
end

\section*{Application}
for \(t=1,2, \cdots, N\)
\[
\begin{aligned}
& u_{t}=-S_{t} x_{t-1} \\
& y_{t} \cdots \operatorname{funct}\left(u_{t}\right)
\end{aligned}
\]
end

\section*{Extended criterion}

The penalty function can be very easily extended to the following form
\[
\left(y_{t}-s_{t}\right)^{2}+\omega u_{t}^{2}+\lambda\left(u_{t}-u_{t-1}\right)^{2}
\]
where the first term leads to the following the output \(y_{t}\) the prescribed set-point \(s_{t}\) and the last term introduces penalization of increments of the control variable. Penalizing the control increments calms control behavior and at the same time it does not result to steady-state deviation of the output and the set-point as it is when penalizing the whole control variable.
\[
\Omega=\left[\begin{array}{ccccccc}
1 & & & & & -s_{t} \\
& \omega+\lambda & & -\lambda & & & \\
& & 0 & & & & \\
& -\lambda & & \lambda & & & \\
& & & & \cdots & & \\
& & & & & 0 & \\
-s_{t} & & & & & & s_{t}^{2}
\end{array}\right]
\]
with \(x_{t}=\left[y_{t}, u_{t}, y_{t-1}, u_{t-1}, \cdots, 1\right]\) the expression \(x_{t}^{\prime} \Omega x_{t}\) gives the extended criterion.

\section*{Control with categorical model}

Model
\[
\begin{gathered}
f\left(y_{t} \mid u_{t}, y_{t-1}\right)=\Theta_{y_{t} \mid u_{t}, y_{t-1}} \\
J=J_{y_{t} \mid u_{t}, y_{t-1}}
\end{gathered}
\]
\begin{tabular}{c|ccccc|cccc}
\(\operatorname{model}(\Theta)\) & & penalty \((J)\) \\
\(u_{t}\) & 1 & 1 & 2 & 2 \\
\(y_{t-1}\) & 1 & 2 & 1 & 2 \\
\hline\(y_{t}=1\) & 0.7 & 0.2 & 0.9 & 0.4 \\
\(y_{t}=2\) & 0.3 & 0.8 & 0.1 & 0.6
\end{tabular} \begin{tabular}{cc|cccc}
\(u_{t}\) & 1 & 1 & 2 & 2 \\
\(y_{t-1}\) & 1 & 2 & 1 & 2 \\
\hline\(y_{t}=1\) & 0 & 1 & 1 & 2 \\
\(y_{t}=2\) & 1 & 2 & 2 & 3
\end{tabular}
where each state is penalized individually. (Above - we do not want big values)

Direct use Bellman equations. Only manipulation is a bit awkward.

\section*{Programs}

\section*{1. T50ctrlMinVar.sce}
minimum variance control
2. T52ctrlDisc.sce
control with categorical model
3. T53ctrlX.sce
control with regression model
4. T54ctrlXEst.sce
adaptive control with regression model

\section*{6 Lecture: Model based classification I}

Clustering: detecting groups (classes) of similar objects creating clusters.
Classification: assigning a new object to one of the existing classes.
Example


Two normal clusters with expectations [1, 2] and [6, 4].
Red cross is a new measurement. It evidently belongs to Cluster 1.

\section*{Generating multimodal data}
- components \(=\) models of individual clusters \(f_{j}\left(x_{t} \mid \theta_{j}\right), j=1,2, \cdots, \nu\)
- pointer \(=\) discrete random process \(c_{t}\) whose values point at the active component

Each cluster has its own model - component.
Example
// Simulation of a mixture with regression components
// ---------------------------------------------
clc, clear, close, mode(0)
```

nd=500; // number of steps
th=[1 5 8]; // component expectations
sd=[1 1 1 1]*.5; // component standard deviations
al=[.3 .4 .3]; // switching probabilities
for t=1:nd
c(t)=sum( cumsum(al)<rand(1,1,'u') ) +1;
x(t)=th(c(t))+sd(c(t))*rand(1,1,'n');
end

```
```

// results
scf();
subplot(211)
plot(x,ones(x),'x', 'markersize',10)
title('Generated values','fontsize',5)
subplot(212)
histplot(20,x);
title('Histogram of generated values','fontsize',5)

```
with the result
Generated values



\section*{Classification 1 - known components}

Given components \(f(x \mid c=i), i=1,2, \cdots \nu\), switching probabilities \(f(c), i=1,2, \cdots \nu\) and one data record \(x=\xi\), estimate the most probable value of \(c\).
\[
f(c \mid x=\xi) \propto f(x=\xi \mid c) f(c)
\]

Example (for \(\nu=3\) )
Components
\[
\begin{aligned}
& f(x \mid c=1)=N_{x}(1,0.5) \\
& f(x \mid c=2)=N_{x}(5,0.5) \\
& f(x \mid c=3)=N_{x}(8,0.5)
\end{aligned}
\]

Model of switching
\[
f(c)=\alpha_{c} \quad \begin{array}{c|ccc}
c & 1 & 2 & 3 \\
\hline \alpha & 0.3 & 0.4 & 0.3
\end{array}
\]

Measurement
\[
x=\xi=2.1
\]

Classification - weights
\[
\begin{gathered}
w_{1} \propto f(c=1 \mid \xi) \propto f(x=2.1 \mid c=1) f(c=1)=\left.N_{x}(1,0.5)\right|_{x=2.1} \alpha_{1}=0.168 \cdot 0.3=0.05 \\
w_{2} \propto 0.00013 \cdot 0.4=0.00005 \\
w_{3} \propto 4 \cdot 10^{-16} \cdot 0.3 \doteq 0
\end{gathered}
\]
- normalization
\[
w=[0.999,0.001,0]
\]
... and we classify to the first class.

\section*{Classification 2 - known pointer for learning \\ \(=\) Learning with a teacher}

Component and pointer models are unknown, values of the pointer are known for learning. \(\rightarrow\) At each step of estimation we update only the component indicated by the pointer.

In practice:

\section*{Learning}

We divide the data sample \(x_{1}, x_{2}, \cdots x_{N}\) into groups \(C_{c}\) with respect to the pointer values \(c=\) \(c_{1}, c_{2}, \cdots c_{N}\) and learn the parameters for all components individually.

\section*{Testing}

Runs as in the previous case.

\section*{Classification 3 - EM-like algorithm}

The expectation-maximization (EM) algorithm is an approach for performing maximum likelihood estimation in the presence of unknown (pointer) variables.

It starts with prior component parameters. Then it repeats the following two steps:
1. determine the values for the pointer variables,
2. estimate the component parameters,
until steady state is reached.
Using the introduced theory, the procedure is like this:
1. Take a dataset \(X=\left[x_{1}, x_{2}, \cdots, x_{N}\right]\) for estimation
2. Set initial components \(f(x \mid c)\) and their stationary probabilities \(\alpha_{c}, c=1,2, \cdots, n_{c}\)
3. Determine weights \(w=f(c \mid X) \propto \alpha_{c} f\left(x_{t} \mid \theta_{c}\right)\) and pointer estimate

For \(t=1: N\)
\[
w_{1}=f(c=1 \mid X) \propto \alpha_{1} f\left(x_{t} \mid \theta_{1}\right)
\]
\[
\begin{gathered}
w_{2}=f(c=2 \mid X) \propto \alpha_{2} f\left(x_{t} \mid \theta_{2}\right) \\
\cdots \\
w_{n_{c}}=f\left(c=n_{c} \mid X\right) \propto \alpha_{n_{c}} f\left(x_{t} \mid \theta_{n_{c}}\right) \\
c_{t}=\arg \max \left(w_{1}, w_{2}, \cdots, w_{n_{c}}\right)
\end{gathered}
\]
4. Recompute component parameters \(\theta\) and switching probabilities \(\alpha\)

For \(j=1: n_{c}\) do
(a) select subset of dataset whose records correspond to pointer value \(j\)
(b) use this subset for estimation of parameters of the \(j\)-th component \(f_{j}\left(x \mid \theta_{j}\right)\) for normal components - average and variance
(c) switching probabilities \(\alpha\) are relative frequencies of the pointer values
5. If the pointer changes go to 3

Remark: It uses learning with a teacher.

\section*{Classification 4-mixture estimation}

Neither model parameters nor pointer values are known. Classification is to be performed with on-line measured data. The procedure is as follows:
1. For each data record \(x_{t}\) determine the weights with respect to currently estimated components \(w_{j}=f\left(c_{t}=j \mid x_{t}\right)\).
2. Data record is added to the statistics with its weight \(S_{j ; t}=S_{j ; t-1}+w_{j}, \kappa_{j ; t}=\kappa_{j ; t-1}+w_{j}\) and point estimates \(\hat{\theta}_{t}\) are computed in a standard way for each component (pointer model can be skipped).

Algorithm
Initial setting: Set initial parameters of components \((\theta, r)\) and corresponding statistics \(S, \kappa\). for \(t=1: n d\)
1. measure data record \(x_{t}\)
2. determine weights \(w\)
for \(j=1: n_{c}\)
(a) \(q_{j}=f\left(x_{t} \mid \theta_{j}\right)\) - proximity
(b) \(w_{j}=\aleph\left(q_{j} \alpha_{j}\right)-\) where \(\aleph\) means normalization to sum equal to 1 end
3. recompute statistics and parameters (e.g. for static normal components) for \(j=1: n_{c}\)
(a) \(S_{j ; t}=S_{j ; t-1}+w_{j} x_{t}\)
(b) \(\kappa_{j ; t}=\kappa_{j ; t-1}+w_{j}\)
(c) \(\gamma_{j ; t}=\gamma_{j ; t-1}=w_{j}\)
(d) \(\theta_{j}=\frac{S_{j ; t}}{\kappa_{j ; t}}\)
(e) \(\alpha_{j}=\aleph(\gamma)\)
end
end

\section*{Remarks}
1. The derivation can be found in the textbook.
2. For component parameters, the point estimates have been used.
3. There are two main points used
(a) pointer estimation for new data record - the basis is \(f(c \mid x)\)
(b) update of statistics with the weight
- standard update: \(S=S+x\)
- for two identical \(x\) and \(x\) it is: \(S=S+2 x\) (weight)
- similarly for \(x\) valid with probability \(w\) it is: \(S=S+w x\) (again weight) and similarly for other statistics.

\section*{Programs}
1. T61classKn.sce
classification with known models of components
2. T62classUnKn.sce
classification with unknown models of components
3. T63EM_C.sce
iterative estimation of pointer and components (like EM algorithm)
4. T64MixReg.sce

Bayesian mixture estimation

\section*{7 Lecture: Model based classification II}

\section*{Naive Bayes}

Estimation of multivariate model can be considerable simplified by the assumption of conditional independence of explanatory variables.

Conditional independence
\[
f\left(x_{1}, x_{2}, \cdots, x_{n} \mid c\right)=\prod_{i=1}^{n} f\left(x_{i} \mid c\right)
\]

Principle of naive Bayes
\[
\begin{gathered}
f(c \mid x) \propto f(x \mid c) f(c)=f\left(x_{1}, x_{2}, \cdots, x_{n} \mid c\right) f(c) \\
=f(c) \prod_{j=1}^{n_{x}} f\left(x_{i} \mid c\right)
\end{gathered}
\]
!! USES ONLY MODELS OF SINGLE VARIABLE !!

KNIME: Task00_NaiveBayes
```

Naive Bayes
Naive Bayes learner and predictor to classify shuttle data.

```


\section*{Logistic regression}

Used for discrete target and continuous explanatory variables.
Starts with Bernoulli model
\[
f\left(c_{t} \mid p\right)=p^{c_{t}}(1-p)^{1-c_{t}}, c_{t}=0,1
\]
\(p=P\left(c_{t}=1\right)\).
Expectation \(E\left[c_{t}\right]=p\) is extended by regression \(b^{\prime} x_{t}=b_{0}+b_{1} x_{1 ; t}+\cdots, b_{m ; t} x_{m}\)
To ensure borders of \(p \in(0,1)\) we model \(\operatorname{logit}(p)=\ln \frac{p}{1-p}\)
\[
\operatorname{logit}(p)=b^{\prime} x_{t}
\]
from which the model is
\[
f\left(c_{t} \mid x_{t}, b\right)=\frac{\exp \left\{c_{t} x_{t} b\right\}}{1+\exp \left\{x_{t} b\right\}}= \begin{cases}\frac{1}{1+\exp \left\{x_{t} b\right\}} & \text { for } c_{t}=0 \\ \frac{\exp \left\{x_{t} b\right\}}{1+\exp \left\{x_{t} b\right\}} & \text { for } c_{t}=1\end{cases}
\]

Usage
\[
z_{t}=b^{\prime} x_{t} \in(-\infty, \infty)
\]
for estimated \(b\) and measured \(x_{t}\) compute \(z_{t}\)
\[
\begin{aligned}
& p=P\left(c_{t}=1\right)=\frac{\exp \left(z_{t}\right)}{1+\exp \left(z_{t}\right)} \\
& \text { for } p>0.5 \text { set } c_{t}=1 \text { else } c_{t}=0
\end{aligned}
\]

\section*{Estimation by ML}
\[
L_{N}(b)=\prod_{t=1}^{N} \frac{\exp \left\{c_{t} x_{t} b\right\}}{1+\exp \left\{x_{t} b\right\}} \rightarrow \ln L_{N}(b)=\sum_{t=1}^{N}\left[c_{t} x_{t} b-\ln \left(1+\exp \left\{x_{t} b\right\}\right)\right]
\]
and maximize numerically.

KNIME: Task01_Logistic_Regression

\section*{Logistic Regression}

Example to building a basic prediction / classification model using logistic regression.

CSV Writer

Logistic


\section*{Poisson regression}

Starts with Poisson model
\[
f\left(c_{t} \mid \lambda\right)=\exp \{-\lambda\} \frac{\lambda^{c_{t}}}{c_{t}!}, c_{t}=0,1,2, \cdots
\]
\(\lambda \geq 0\) is intensity of occurring events.
To ensure nonnegativity of \(\lambda\), we extend \(\ln (\lambda)=b^{\prime} x_{t}=b_{0}+b_{1} x_{1 ; t}+\cdots b_{m} x_{m ; t}\)
\[
\rightarrow \lambda=\exp \left(b^{\prime} x_{t}\right)
\]

Model in logarithm
\[
\ln \left(f\left(c_{t} \mid b, x_{t}\right)\right)=-\exp \left\{x_{t} b\right\}+c_{t} x_{t} b-\ln \left(c_{t}!\right)
\]

\section*{Estimation by LN}

Log-likelihood is
\[
\ln L_{N}(b)=\sum_{t=1}^{N}\left[-\exp \left\{x_{t} b\right\}+c_{t} x_{t} b-\ln \left(c_{t}!\right)\right]
\]
and it is maximized numerically.

\section*{8 Lecture: Clustering}

We have multimodal data \(x\) and want to capture density clusters.

\section*{K-means clustering}

0 . Set \(n\) initial cluster centers ( \(n\) fixed)
1. To each data point \(x_{i}\) assign the nearest center.

The assigned points to a center form the cluster.
2. For each cluster compute its centroid (point average)
3. Shift the centers to the centroids
4. Repeat from 1 if changes occur

Example


KNIME: Task02_k-Means_Clustering


\section*{K-medoids clustering}

Similar to \(k\)-means.

0 . Determine \(m d\) as the desired number of clusters. Randomly select \(m d\) data points as initial centers of medoids.

0 . To each medoid find the points that are closest to it. They will be initial clusters.
0 . Determine overall distance of points from their medians.
1. Randomly select one medoid and one non-medoid (data point that is not a medoid).
2. Swap them and again determine overall distance of points from their medians.
3. If the distance is smaller, continue by 1 . If not, algorithm ends.

KNIME: Task03_k-Medoids_Clustering


\section*{Fuzzy clustering (c-means)}

In the c-means algorithm we minimize criterion
\[
J=\sum_{i=1}^{N} \sum_{j=1}^{C} u_{i j}^{m}\left\|x_{i}-c_{j}\right\|^{2}, m \geq 1
\]
where \(u_{i j}\) is a degree of membership of the point \(x_{i}\) to cluster \(c_{j}\) and \(\|\cdot\|\) is a norm.
The update of weights \(u_{i j}\) is performed as follows
- determine the centers (weighted average - follows from the criterion)
\[
c_{j}=\frac{\sum_{i=1}^{N} u_{i j}^{m} x_{i}}{\sum_{i=1}^{N} u_{i j}^{m}}
\]
- weights (are given as membership functions)
\[
\begin{equation*}
u_{i j}=\frac{1}{\sum_{k=1}^{C}\left(\frac{\left\|x_{i}-c_{j}\right\|}{\left\|x_{i}-c_{k}\right\|}\right)^{\frac{2}{m-1}}} \tag{3}
\end{equation*}
\]

Algorithm
0 . Set the initial matrix of membership \(U\).
1. Compute the centers \(c_{j}\) with existing matrix \(U\).
2. Update the matrix \(U\).
3. If \(\left\|U_{\text {nová }}-U_{\text {staráa }}\right\|<\epsilon\), END otherwise go to 1 .

KNIME: Task04_c-Means_Clustering

\section*{c-Means Clustering}

Fuzzy clustering of the iris dataset


\section*{Density based clustering (dbscan)}

We have a set of data \(X=\left\{x_{1}, x_{2}, \cdots, x_{N}\right\}\), where \(x_{i} \in R^{m}\)
We define:
- Distance of two points \(x\) and \(y\) and denote it by \(d(x, y)\).
- \(\epsilon\)-neighborhood of point \(x\)
\[
O_{\epsilon}(x)=\{x \in X: d(x, y)<\epsilon\} .
\]
- Inner point is such one that has in its neighborhood at least given number of points.
- A point \(y\) is accessible from the point \(x\), if a sequence of inner points from \(x\) to \(y\) exists.
- A connection between points \(x\) a \(y\) exists, it both these points are accessible from some inner point.


Algorithm of clustering
1. For each point from \(X\) find its \(\epsilon\)-neighborhood.
2. Define variables "clus" and "buff" (for storing points).
3. To "clus" put a single inner point and to "buff" its neighborhood.
4. Select one point (e.g. the first one) from "buff". Add it to "cluss" and its neighborhood add to "buff".
5. From "buff" remove all points that have already been used (those that are in some cluster).
6. Repeat from 4. until "buff" is not empty. Otherwise continue.
7. Remember the created cluster "clus" and prepare the variable for new one.
8. If there exists another free inner point, put it to "clus" and go to 4 . If not, stop the algorithm.

Clusters are formed by points that are connected.

KNIME: Task05_Density _Clustering


\section*{Hierarchical clustering (agglomerative)}
1. All data points are denoted as clusters on the level 1 (with only one point).
2. Find two nearest clusters and join them together in a new point. Its level is equal to the number of points in joined in this new point.
3. The coordinates of the cluster lie on a connecting line of the coordinates of clusters to be joined in the proportion of their levels (the higher level the nearer).
4. Remember the clusters from which the new one has been created (hierarchy).
5. Repeat from 2 until only one cluster remains.

For more information and the divisive version of the algorithm see the textbook.

\section*{Example}

The data are \(\quad x=[1.4,1.8,2.5,4.2,4.7,6.5]\).
Construct dendrogram.


KNIME: Task06_Hierarchical_Clustering


\section*{9 Lecture: Classification}

\section*{K-nearest neighbour}

We have data \(X=\left\{x_{i}\right\}_{i=1}^{N}\) with detected clusters. The task is: assign a newly measured point \(y\) to some cluster.

Algorithm
1. Compute the distance of the point \(y\) from all points from \(x_{i} \in X\).
2. Determine \(k\) points \(x_{i}, i=1,2, \cdots, k\) nearest to \(y\).
3. Assign \(y\) to the cluster to which majority of the \(k\) nearest points belongs.

KNIME: Task07_k-NearNeighb


\section*{Decision trees}

We have discrete data records \(x_{t}=\left[x_{1}, x_{2}, \cdots, x_{n}\right]_{t}, t=1,2, \cdots, N\) and a pointer variable \(c_{t} \in\) \(\{1,2, \cdots, m\}\) which assigns the data records \(x_{t}\) to one of \(m\) classes.

\section*{Example}

Let us have the following data
\begin{tabular}{c|cc|c}
\(t\) & \(x_{1}\) & \(x_{2}\) & \(c\) \\
\hline 1 & 1 & 1 & 1 \\
2 & 1 & 2 & 1 \\
3 & 2 & 1 & 2 \\
4 & 2 & 2 & 2
\end{tabular}
where \(x_{1}, x_{2}\) are data records and \(c\) is pointer variable.
We chose the root cluster as \(x_{1}\) with values \(\{1,2\}\). Then,
- if \(x_{1}=1\) then \(x_{2} \in\{1,2\}\)
- if \(x_{1}=1\) and \(x_{2}=1\) then \(c=1\)
- if \(x_{1}=1\) and \(x_{2}=2\) then \(c=1\)
- if \(x_{1}=2\) then \(x_{2} \in\{1,2\}\)
- if \(x_{1}=2\) and \(x_{2}=1\) then \(c=2\)
- if \(x_{1}=2\) and \(x_{2}=2\) then \(c=2\)


Now, we measure \(x_{t}=[1,2]\). Using the tree, we classify it to \(c_{t}=1\)
Problem: What order of the variables in the tree is the best one.

KNIME: Task08_Decision_Tree

\section*{Decision Tree}

Prediction (classification) using a decision tree. Dataset describes wine chemical features.
Output class is wine color: red/white


\section*{Support vector machines}

We have a sequence of data points \(x_{i}, i=1,2, \cdots, n\). Some of them have the attribute + and the rest - . We are to separate them so that the distance of the line (hyperplane) from the + points and -points would be maximal.


Let us denote the separating line as \(y=\alpha y+\beta=0\). Than, we look for maximal \(\delta\) such that two parallel lines \(y=\alpha y+\beta+\delta=0\) and \(y=\alpha y+\beta-\delta=0\) also separate the points - i.e. the points are separated by a strip of the width \(2 \delta\).

The task leads to numerical optimization of nonlinear function.

KNIME: Task09_Support_Vec_Mach
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